

## 15.10 Exercises: Some Solutions (November 26, 2011)

15.1. Prove Proposition 15.5 (p. 340).

• SOLUTION TO EX. 15.1:

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15.2. Prove Proposition 15.12 (p. 351).

15.3. Prove (see page 352):

$$\tilde{\boldsymbol{\beta}} - \tilde{\boldsymbol{\beta}}_{(i)} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{i}_i \frac{\hat{\varepsilon}_i}{\hat{m}_{ii}}, \quad \frac{\hat{\varepsilon}_i}{\hat{m}_{ii}} = \tilde{\delta} = \text{BLUE}(\delta \mid \mathcal{M}_Z).$$

15.4. Consider the models defined as in (15.144) (p. 351). Show that the  $F$ -test statistics for testing  $\delta = 0$  under  $\mathcal{M}_Z$  is

$$t_i^2(\mathbf{V}) = \frac{\hat{\varepsilon}_i^2}{\hat{m}_{ii}\hat{\sigma}_{(i)}^2} = \frac{\tilde{\delta}_i^2}{\widetilde{\text{var}}(\tilde{\delta}_i)},$$

where  $\hat{\varepsilon} = \hat{\mathbf{M}}\mathbf{y} = \mathbf{V}^{-1}\tilde{\boldsymbol{\varepsilon}}$ ,  $\tilde{\boldsymbol{\varepsilon}}$  is the BLUE's residual, and  $(n-p-1)\hat{\sigma}_{(i)}^2 = \text{SSE}_{(i)}(\mathbf{V}) = \text{SSE}_Z(\mathbf{V})$ . The statistics  $t_i^2(\mathbf{V})$  is a generalized version of the externally Studentized residual, cf. (8.192) (p. 182).

15.5. Suppose that  $\mathbf{V}$  has the intraclass correlation structure and  $\mathbf{1} \in \mathcal{C}(\mathbf{X})$ , where  $\mathbf{X}_{n \times p}$  has full column rank. Show that then  $\text{DFBETA}_i(\mathbf{V}) = \text{DFBETA}_i(\mathbf{I})$ . What about the equality between Cook's distance  $\text{COOK}_i^2$  and the generalized Cook's distance  $\text{COOK}_i^2(\mathbf{V})$ ?

15.6. Express the Cook's distance  $\text{COOK}_i^2(\mathbf{V})$  as a function of an appropriate Mahalanobis distance.

15.7. Consider a partitioned model  $\mathcal{M}_{12} = \{\mathbf{y}, \mathbf{X}\boldsymbol{\beta}, \mathbf{V}\}$ , where  $\mathbf{X}$  has full column rank and  $\mathbf{V}$  is positive definite. Suppose that  $\hat{\boldsymbol{\beta}}_1$  is fully efficient under the small model  $\mathcal{M}_1 = \{\mathbf{y}, \mathbf{X}_1\boldsymbol{\beta}_1, \mathbf{V}\}$ . Show that

$$\text{eff}(\hat{\boldsymbol{\beta}}_1 \mid \mathcal{M}_{12}) = 1 \iff \mathbf{X}'_1\mathbf{X}_2\tilde{\boldsymbol{\beta}}_2(\mathcal{M}_{12}) = \mathbf{X}'_1\mathbf{X}_2\hat{\boldsymbol{\beta}}_2(\mathcal{M}_{12}).$$

Isotalo, Puntanen & Styan (2007).

15.8. Consider the models  $\mathcal{M}$  and  $\mathcal{M}_{(i)}$ , defined in (15.144) (p. 351), where  $\mathbf{X}$  and  $(\mathbf{X} : \mathbf{i}_i)$  have full column ranks. Suppose that  $\hat{\boldsymbol{\beta}} = \tilde{\boldsymbol{\beta}}$  under  $\mathcal{M}$ .

Show that the equality  $\hat{\beta}_{(i)} = \tilde{\beta}_{(i)}$  for all  $i = 1, \dots, n$  holds if and only if  $\mathbf{MVM} = c^2\mathbf{M}$  for some nonzero  $c \in \mathbb{R}$ .

Nurhonen & Puntanen (1992a), Isotalo, Puntanen & Styan (2007).

**15.9.** According to (8.192) (p. 182), the  $F$ -test statistic for testing the hypothesis  $\delta = 0$  under the model  $\mathcal{M}_Z(\mathbf{I}) = \{\mathbf{y}, \mathbf{X}\boldsymbol{\beta} + \mathbf{u}\delta, \sigma^2\mathbf{I}\}$ , where  $\mathbf{u} = \mathbf{i}_i$ , becomes

$$t_i^2 = \frac{\mathbf{y}'\mathbf{P}_{\mathbf{Mu}}\mathbf{y}}{\frac{1}{n-p-1}\mathbf{y}'(\mathbf{M} - \mathbf{P}_{\mathbf{Mu}})\mathbf{y}} = \frac{\hat{\varepsilon}_i^2}{\frac{1}{n-p-1} \text{SSE}_{(i)} m_{ii}},$$

which is the squared externally Studentized residual. Under the model  $\mathcal{M}_Z(\mathbf{I})$  the test statistic  $t_i^2$  follows an  $F$ -distribution with 1 and  $n - p - 1$  degrees of freedom. Denote  $\mathcal{M}_Z(\mathbf{V}) = \{\mathbf{y}, \mathbf{X}\boldsymbol{\beta} + \mathbf{u}\delta, \sigma^2\mathbf{V}\}$ . Prove that under the model  $\mathcal{M}_Z(\mathbf{V})$ :  $t_i^2 \sim F(1, n - p - 1) \iff \mathbf{MVM} = c^2\mathbf{M}$  for some  $c \neq 0$ .

Nurhonen & Puntanen (1991), Rao & Mitra (1971b, Ch. 9).

**15.10.** Write up the solution to Exercise 15.9 when  $\mathbf{Z} = (\mathbf{1} : \mathbf{i}_n)$ . Moreover, confirm that the following statements are equivalent:

- (a)  $\mathbf{CVC} = c^2\mathbf{C}$  for some  $c \neq 0$ ,
- (b)  $\mathbf{V} = \alpha^2\mathbf{I} + \mathbf{a}\mathbf{1}' + \mathbf{1}\mathbf{a}'$ , where  $\mathbf{a}$  is an arbitrary vector and  $\alpha$  is any scalar ensuring the positive definiteness of  $\mathbf{V}$ .

Above  $\mathbf{C}$  denotes the centering matrix. Confirm also that the eigenvalues of  $\mathbf{V}$  in (b) are  $\alpha^2 + \mathbf{1}'\mathbf{a} \pm \sqrt{n\mathbf{a}'\mathbf{a}}$ , each with multiplicity one, and  $\alpha^2$  with multiplicity  $n - 2$ . These results appear useful when studying the robustness of the Grubbs's test for detecting a univariate outlier.

Baksalary & Puntanen (1990b), Baksalary, Nurhonen & Puntanen (1992), Lehman & Young (1993), Markiewicz (2001).

**• SOLUTION TO EX. 15.10:**

Consider first the equation

$$\mathbf{CVC} = c^2\mathbf{C}, \quad \text{where } c \neq 0. \quad (1)$$

Our task is to characterize all positive definite matrices  $\mathbf{V}$  which satisfy (1). Baksalary & Puntanen (1990b) proceed by using Theorem 2 of Baksalary (1984), which gives the following general positive definite solution:

$$\mathbf{V} = c^2\mathbf{C} + \mathbf{T}_1\mathbf{w}\mathbf{1}' + \mathbf{1}\mathbf{w}'\mathbf{T}_1' + \nu\mathbf{1}\mathbf{1}', \quad (2)$$

where  $\nu$  is an arbitrary scalar and  $\mathbf{w} \in \mathbb{R}^{n-1}$  is an arbitrary vector satisfying the condition

$$\mathbf{w}'\mathbf{w} < c^2\nu.$$

Here, as in the solution of Exercise 3.10 (p. 104),

$$\mathbf{T} = (\mathbf{t}_1 : \dots : \mathbf{t}_{n-1} : \mathbf{t}_n) = (\mathbf{T}_1 : \frac{1}{\sqrt{n}}\mathbf{1}_n),$$

where the vectors  $\mathbf{t}_i$  are the orthonormal eigenvectors of  $\mathbf{C}$  and  $\mathbf{C}$  has the eigenvalue decomposition

$$\mathbf{C} = \mathbf{T}\mathbf{\Lambda}\mathbf{T}' = \mathbf{T} \begin{pmatrix} \mathbf{I}_{n-1} & \mathbf{0} \\ \mathbf{0}' & 0 \end{pmatrix} \mathbf{T}' = \mathbf{T}_1\mathbf{T}_1'.$$

It is interesting to rewrite (1) as

$$\mathbf{C}(c^{-2}\mathbf{V})\mathbf{C} = \mathbf{C}, \quad \text{i.e.,} \quad c^{-2}\mathbf{V} \in \{\mathbf{C}_{\text{pd}}^-\}.$$

In view of the solution to Exercise 3.10, the general expression for  $c^{-2}\mathbf{V} \in \{\mathbf{C}_{\text{pd}}^-\}$  is

$$c^{-2}\mathbf{V} = \mathbf{C} + \mathbf{1}\mathbf{f}'\mathbf{T}_1' + \mathbf{T}_1\mathbf{f}\mathbf{1}' + \delta\mathbf{1}\mathbf{1}', \quad (3)$$

for any  $\mathbf{f} \in \mathbb{R}^{n-1}$  and  $\delta \in \mathbb{R}$  which satisfy

$$\mathbf{f}'\mathbf{f} < \delta. \quad (4)$$

Multiplying (3) by  $c^2$  gives the general pd expression for  $\mathbf{V}$ :

$$\begin{aligned} \mathbf{V} &= c^2\mathbf{C} + c^2\mathbf{1}\mathbf{f}'\mathbf{T}_1' + c^2\mathbf{T}_1\mathbf{f}\mathbf{1}' + c^2\delta\mathbf{1}\mathbf{1}' \\ &:= c^2\mathbf{C} + \mathbf{1}\mathbf{w}'\mathbf{T}_1' + \mathbf{T}_1\mathbf{w}\mathbf{1}' + c^2\delta\mathbf{1}\mathbf{1}', \end{aligned} \quad (5)$$

which is of the form (2).

Baksalary & Puntanen (1990b) show that the sets specified by (b) and (2) are equal. Let's copy part their proof here.

Because the columns of  $(\mathbf{T}_1 : \mathbf{1})$  span the entire  $\mathbb{R}^n$ , it is clear that  $\mathbf{a}$  in (b) can be uniquely decomposed as  $\mathbf{a} = \mathbf{T}_1\mathbf{a}_1 + \gamma\mathbf{1}$  for some  $\mathbf{a}_1 \in \mathbb{R}^{n-1}$  and  $\gamma \in \mathbb{R}$ , and hence

$$\begin{aligned} \mathbf{T}_1\mathbf{a}_1 &= \mathbf{a} - \gamma\mathbf{1}, \\ \mathbf{T}_1\mathbf{a}_1\mathbf{1}' &= \mathbf{a}\mathbf{1}' - \gamma\mathbf{1}\mathbf{1}', \quad \mathbf{1}\mathbf{a}_1'\mathbf{T}_1' = \mathbf{1}\mathbf{a}' - \gamma\mathbf{1}\mathbf{1}'. \end{aligned}$$

Choosing  $\mathbf{w} = \mathbf{a}_1$  in (b) yields

$$\begin{aligned} \mathbf{V} &= c^2\mathbf{C} + \mathbf{a}\mathbf{1}' + \mathbf{1}\mathbf{a}' - 2\gamma\mathbf{1}\mathbf{1}' + \nu\mathbf{1}\mathbf{1}' \\ &= c^2\mathbf{C} + \mathbf{a}\mathbf{1}' + \mathbf{1}\mathbf{a}' + (\nu - 2\gamma)\mathbf{1}\mathbf{1}' \\ &= c^2(\mathbf{I} - \mathbf{J}) + \mathbf{a}\mathbf{1}' + \mathbf{1}\mathbf{a}' + n(\nu - 2\gamma)\mathbf{J}. \end{aligned}$$

Substituting then  $c^2 = \alpha^2$  and  $\nu = 2\gamma + \alpha^2/n$  into (2) yields (b). This shows that the class (2) includes all matrices having the structure (b); (b) is called

the class of *Baldessari's structure*. The converse inclusion follows similarly, by substituting  $\alpha^2 = c^2$  and  $\mathbf{a} = \mathbf{T}_1 \mathbf{a}_1 + \gamma \mathbf{1}$ , where  $\gamma = \nu/2 - c^2/(2n)$ .

How to confirm that the eigenvalues of

$$\mathbf{V} = \alpha^2 \mathbf{I} + \mathbf{a} \mathbf{1}' + \mathbf{1} \mathbf{a}' := \alpha^2 \mathbf{I} + \mathbf{W}$$

are  $\alpha^2 + \mathbf{1}' \mathbf{a} \pm \sqrt{n \mathbf{a}' \mathbf{a}}$ , each with multiplicity one, and  $\alpha^2$  with multiplicity  $n - 2$ ? We observe that

$$\text{ch}(\mathbf{V}) = \{\alpha^2 + \text{ch}(\mathbf{a} \mathbf{1}' + \mathbf{1} \mathbf{a}')\}. \quad (\text{c})$$

If  $\text{rk}(\mathbf{1} : \mathbf{a}) = 2$ , then  $\mathbf{q} \in \mathcal{C}(\mathbf{1} : \mathbf{a})^\perp$  as stated in the solution of Exercise 3.10, and thereby 0 is an eigenvalue of  $\mathbf{W}$  of multiplicity 2. The remaining two eigenvalues of  $\mathbf{W} = \mathbf{a} \mathbf{1}' + \mathbf{1} \mathbf{a}'$  are  $\mathbf{1}' \mathbf{a} \pm \sqrt{n \mathbf{a}' \mathbf{a}}$ . We skip the proof.

Some references; those not appearing in the Tricks References, written in full:

- Chaganty, N. Rao & Vaish, A. K. (1997). An invariance property of common statistical tests. *Linear Algebra and its Applications*, 264, 421–437.
- Farebrother, R. W. (1987). Three theorems with applications to Euclidean distance matrices. *Linear Algebra and its Applications*, 95, 11–16.
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- Mathew (1985).
- Sharpe, G. E. & Styan, G. P. H. (1965). Circuit duality and the general network inverse. *IEEE Trans. Circuit Theory CT-12*, 22–27.
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